

# Algorithms for Data Science: The EM Algorithm

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Instead of the combinatorial approach of the K-means algorithm, take a more direct probabilistic approach to modeling distribution P(X).

Assume each of the K clusters corresponds to a multivariate distribution  $P_k(X)$ ,

P(X) is then a *mixture* of these distributions as  $P(X) = \sum_{k=1}^{K} \pi_k P_k(X).$ 

Specifically, take  $P_k(X)$  as a multivariate normal distribution  $f_k(X) = N(\mu_k, \sigma_k^2 I)$ 

and mixture density  $f(X) = \sum_{k=1}^{K} \pi_k f_k(X)$ .

Use Maximum Likelihood to estimate parameters

$$heta = (\mu_1, \ldots, \mu_K, \sigma_1^2, \ldots, \sigma_K^2, \pi_1, \ldots, \pi_K)$$

based on their log-likelihood

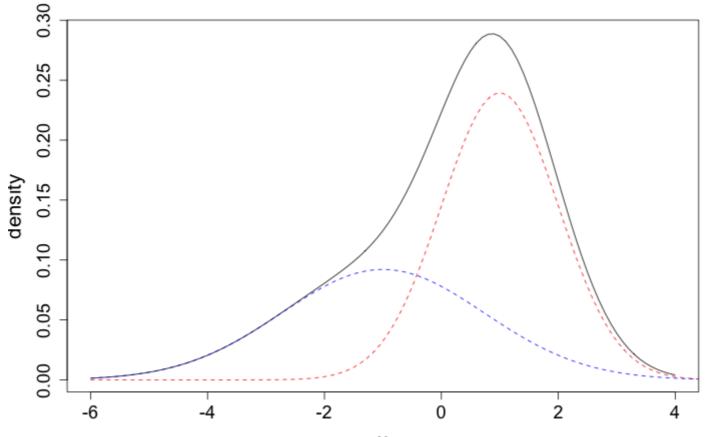
$$\ell( heta;X) = \sum_{i=1}^N \log \left[\sum_{k=1}^K \pi_k f_k(x_i; heta)
ight]$$

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Maximizing this likelihood directly is computationally difficult

Use Expectation Maximization algorithm (EM) instead.

#### Example: Mixture of Two Univariate Gaussians



Consider unobserved latent variables  $\Delta_{ik}$  taking values 0 or 1,

 $\Delta_{ij} = 1$  specifies observation  $x_i$  was generated by component k of the mixture distribution.

Now set  $Pr(\Delta_{ik} = 1) = \pi_k$ , and assume we *observed* values for latent variables  $\Delta_{ik}$ .

We can write the log-likelihood in this case as

$$\ell_0( heta;X,\Delta) = \sum_{i=1}^N \sum_{k=1}^K \Delta_{ik} \log f_k(x_i; heta) + \sum_{i=1}^N \sum_{k=1}^K \Delta_{ik} \log \pi_k$$

We have closed-form solutions for maximum likelihood estimates:

$$\hat{\mu}_k = rac{\sum_{i=1}^N \Delta_{ik} x_i}{\sum_{i=1}^N \Delta_{ik}} \ \hat{\sigma}_k^2 = rac{\sum_{i=1}^N \Delta_{ik} (x_i - \hat{\mu}_k)^2}{\sum_{i=1}^N \Delta_{ik}}$$

$$\hat{\pi}_k = rac{\sum_{i=1}\Delta_{ik}}{N}.$$

We have a problem of type

$$egin{array}{lll} \min_x & f_0(x) \ ext{s.t.} & f_i(x) \leq 0 \; i=1,\ldots,m \ & h_i(x)=0 \; i=1,\ldots,p \end{array}$$

Note: This discussion follows Boyd and Vandenberghe, *Convex Optimization* 

To solve these type of problems we will look at the *Lagrangian* function:

$$L(x,\lambda,
u)=f_0(x)+\sum_{i=1}^m\lambda_if_i(x)+\sum_{i=1}^p
u_ig_i(x)$$

There is a beautiful result giving *optimality conditions* based on the Lagrangian:

Suppose  $\tilde{x}$ ,  $\tilde{\lambda}$  and  $\tilde{\nu}$  are *optimal*, then

$$egin{aligned} f_i( ilde{x}) &\leq 0 \ h_i( ilde{x}) &= 0 \ & ilde{\lambda_i} &\geq 0 \ & ilde{\lambda_i} f_i( ilde{x}) &= 0 \ & ilde{
aligned} 
onumber \ &
abla L( ilde{x}, ilde{\lambda}, ilde{
u}) &= 0 \end{aligned}$$

We can use the gradient and feasibility conditions to prove the MLE result.

Of course, this result depends on observing values for  $\Delta_{ik}$  which we don't observe. Use an iterative approach as well:

- given current estimate of parameters  $\theta$ ,
- Substitute  $E[\Delta_{ik}|X_i, \theta]$  for  $\Delta_{ik}$ .

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We will prove that this maximizes the likelihood we need  $\ell(\theta; X)$ .

## Soft K-means Clustering

In the mixture case, what does this look like?

Define

$$\gamma_{ik}( heta) = E(\Delta_{ik}|X_i, heta) = Pr(\Delta_{ik}=1|X_i, heta)$$

### Soft K-means Clustering

Use Bayes' Rule to write this in terms of the multivariate normal densities with respect to current estimates  $\theta$ :

$$egin{aligned} &\gamma_{ik} = rac{Pr(X_i | \Delta_{ik} = 1) Pr(\Delta_{ik} = 1)}{Pr(X_i)} \ &= rac{f_k(x_i; \mu_k, \sigma_k^2) \pi_k}{\sum_{l=1}^K f_l(x_i; \mu_l, \sigma_l^2) \pi_l} \end{aligned}$$

# Soft K-means Clustering

Quantity  $\gamma_{ik}(\theta)$  is referred to as the *responsibility* of cluster k for observation i, according to current parameter estimate  $\theta$ .

## Soft K-means Clustering

We can now give a complete specification of the EM algorithm for mixture model clustering.

- 1. Take initial guesses for parameters heta
- 2. *Expectation Step*: Compute responsibilities  $\gamma_{ik}(\theta)$
- 3. *Maximization Step*: Estimate new parameters based on responsibilities as below.
- 4. Iterate steps 2 and 3 until convergence

### Soft K-means Algorithm

Estimates in the Maximization step are given by

$$\hat{\mu}_k = rac{\sum_{i=1}^N \gamma_{ik}( heta) x_i}{\sum_{i=1}^N \gamma_{ik}}$$

$$\hat{\sigma}_k^2 = rac{\sum_{i=1}^N \gamma_{ik}( heta)(x_i-\mu_k)^2}{\sum_{i=1}^N \gamma_{ik}( heta)}$$

and

 $\neg N$ ( )

## Soft K-means Algorithm

The name "soft" K-means refers to the fact that parameter estimates for each cluster are obtained by weighted averages across all observations.

So, why does that work?

Why does plugging in  $\gamma_{ik}(\theta)$  for the latent variables  $\Delta_{ik}$  work?

Why does that maximize log-likelihood  $\ell(\theta; X)$ ?

Think of it as follows:

Z: observed data  $Z^m:$  missing  $\mathit{latent}$  data  $T=(Z,Z^m):$  complete data (observed and missing)

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 $\ell(\theta'; Z)$ : log-likehood w.r.t. *observed* data  $\ell_0(\theta'; T)$ : log-likelihood w.r.t. *complete* data

Next, notice that

$$Pr(Z| heta') = rac{Pr(T| heta')}{Pr(Z^m|Z, heta')}$$

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As likelihood:

$$\ell( heta';Z) = \ell_0( heta';T) - \ell_1( heta';Z^m|Z)$$

Iterative approach: given parameters  $\boldsymbol{\theta}$  take expectation of log-likelihoods

$$egin{aligned} \ell( heta';Z) &= & E[\ell_0( heta';T)|Z, heta] - E[\ell_1( heta';Z^m|Z)|Z, heta] \ &\equiv & Q( heta', heta) - R( heta', heta) \end{aligned}$$

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In soft k-means,  $Q(\theta', \theta)$  is the log likelihood of complete data with  $\Delta_{ik}$  replaced by  $\gamma_{ik}(\theta)$ 

The general EM algorithm

- 1. Initialize parameters  $heta^{(0)}$
- 2. Construct *function*  $Q(\theta', \theta^{(j)})$
- 3. Find next set of parameters  $heta^{(j+1)} = rg\max_{ heta'} Q( heta', heta^{(j)})$
- 4. Iterate steps 2 and 3 until convergence

So, why does that work?

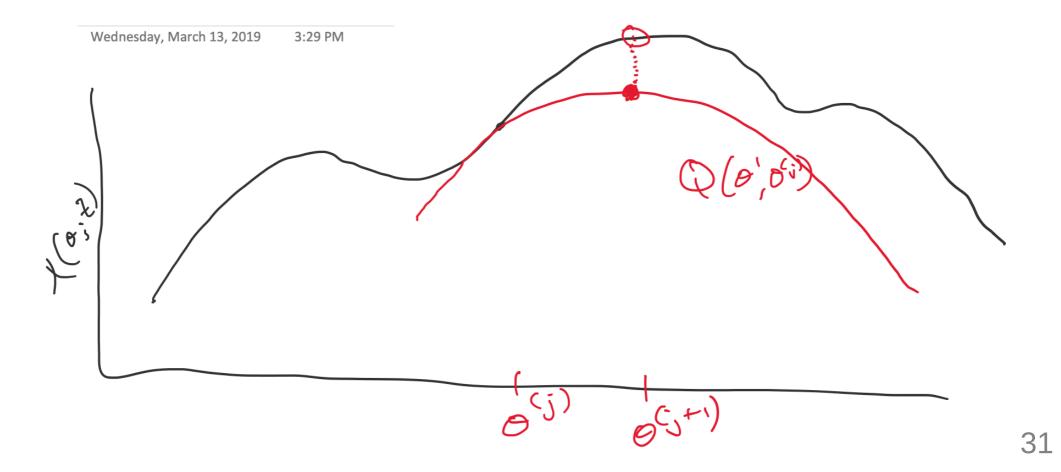
$$\ell( heta^{(j+1)};Z) - \ell( heta^{(j)};Z) = egin{array}{c} [Q( heta^{(j+1)}, heta^{(j)}) - Q( heta^{(j)}, heta^{(j)})] \ - [R( heta^{(j+1)}, heta^{(j)}) - R( heta^{(j)}, heta^{(j)})] \ \geq egin{array}{c} 0 \end{array}$$

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I.E., every step makes log-likehood larger

#### Why else does it work? Q( heta', heta) minorizes $\ell( heta';Z)$



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General algorithmic concept:

Iterative approach:

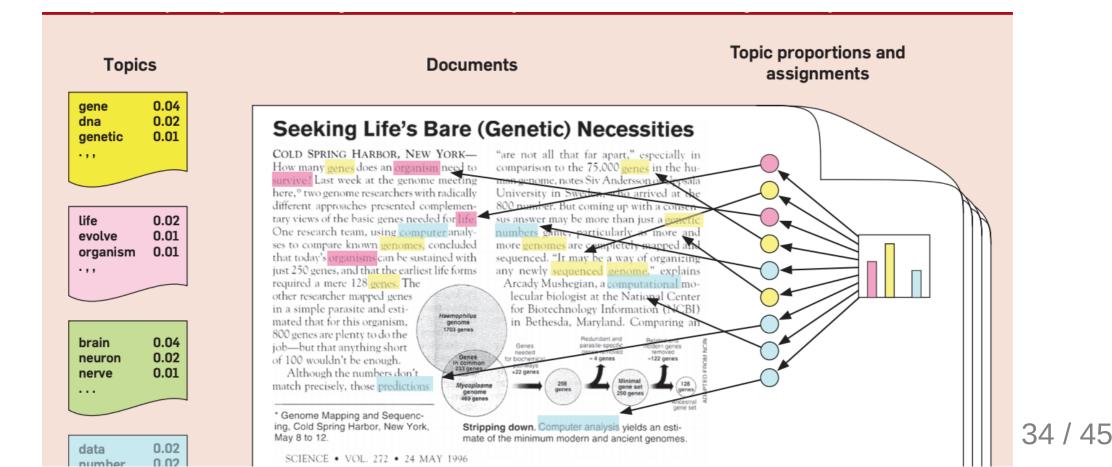
- Initialize parameters
- Construct bound based on current parameters
- Optimize bound

#### Imputing missing data

Z: observed data  $Z^m$ : missing observations

Requires a likelihood model...

#### Documents as *mixtures* of topics (Hoffman 1998)



We have a set of documents D

Each document modeled as a bag-of-words (bow) over dictionary W.

 $x_{w,d}$ : the number of times word  $w \in W$  appears in document  $d \in D$ .

Let's start with a simple model based on the frequency of word occurrences.

Each document is modeled as  $n_d$  draws from a *Multinomial* distribution with parameters  $\theta_d = \{\theta_{1,d}, \dots, \theta_{W,d}\}$ 

Note  $heta_{w,d} \geq 0$  and  $\sum_w heta_{w,d} = 1$ .

Probability of observed corpus D

$$Pr(D|\{ heta_d\}) \propto \prod_{d=1}^D \prod_{w=1}^W heta_{w,d}^{x_{w,d}}$$

Problem 1:

Prove MLE 
$$\hat{ heta}_{w,d} = rac{x_{w,d}}{n_d}$$

Let's change our document model to introduce topics.

The key idea is that the probability of observing a *word* in a *document* is given by two pieces:

- The probability of observing a *topic* in a document, and
- The probability of observing a *word* given a *topic*

$$Pr(w,d) = \sum_{t=1}^T Pr(w|t) Pr(t|d)$$

So, we rewrite corpus probability as

$$Pr(D|\{p_d\}\{ heta_t\}) \propto \prod_{d=1}^D \prod_{w=1}^W \left(\sum_{t=1}^T p_{t,d} heta_{w,t}
ight)^{x_{w,d}}$$

So, we rewrite corpus probability as

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ight)^{x_{w,d}}$$

Mixture of topics!!

#### A fully observed model

Assume you know the *latent* number of occurences of word w in document d generated from topic t:

$$\Delta_{w,d,t}$$
 , such that  $\sum_t \Delta_{w,d,t} = x_{w,d}$  .

In that case we can rewrite corpus probability:

$$Pr(D|\{p_d\},\{ heta_t\}) \propto \prod_{d=1}^D \prod_{w=1}^W \prod_{t=1}^T (p_{t,d} heta_{w,t})^{\Delta_{w,d,t}}$$

**Problem 2** Show MLEs given by

$$\begin{split} \hat{p}_{t,d} &= \frac{\sum_{w=1}^{W} \Delta_{w,d,t}}{\sum_{t=1}^{T} \sum_{w=1}^{W} \Delta_{w,d,t}} \\ \hat{\theta}_{w,t} &= \frac{\sum_{d=1}^{D} \Delta_{w,d,t}}{\sum_{w=1}^{W} \sum_{d=1}^{D} \Delta_{w,d,t}} \end{split}$$

Since we don't observe  $\Delta_{w,d,t}$  we use the EM algorithm

At each iteration (given current parameters  $\{p_d\}$  and  $\{\theta_d\}$  find *responsibility* 

$$\gamma_{w,d,t} = E[\Delta_{w,d,t}|\{p_d\},\{ heta_t\}]$$

and maximize fully observed likelihood plugging in  $\gamma_{w,d,t}$  for  $\Delta_{w,d,t}$ 

Problem 4: Show

$$\gamma_{w,d,t} = x_{w,d} imes rac{p_{t,d} heta_{w,t}}{\sum_{t'=1}^T p_{t',d} heta_{w,t'}}$$