

# Introduction to Data Science: Statistical Principles

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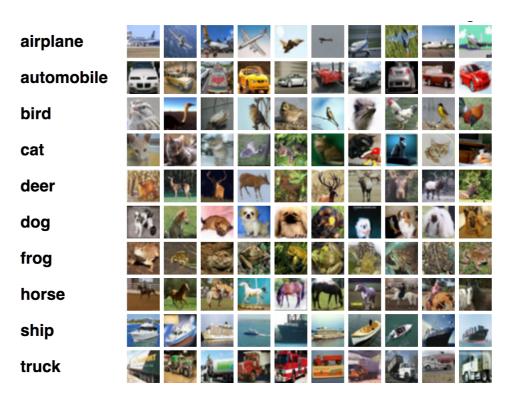
In this class we learn *Statistical and Machine Learning* techniques for data analysis.

By the time we are done, you should

- be able to read critically papers or reports that use these methods.
- be able to use these methods for daata analysis

In either case, you will need to ask yourself if findings are **statistically significant**.

- Use a classification algorithm to distinguish images
- Accurate 70 out of 100 cases.
- Could this happen by chance alone?



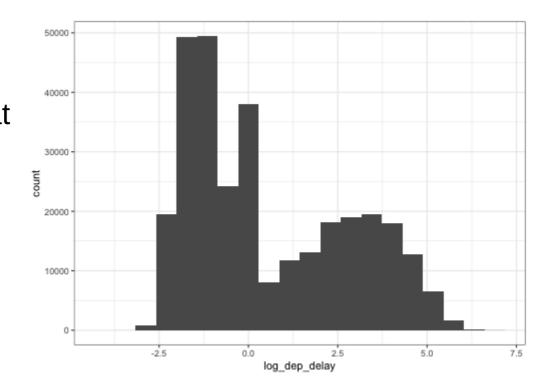
To be able to answer these question, we need to understand some basic probabilistic and statistical principles.

In this course unit we will review some of these principles.

### Variation, randomness and stochasticity

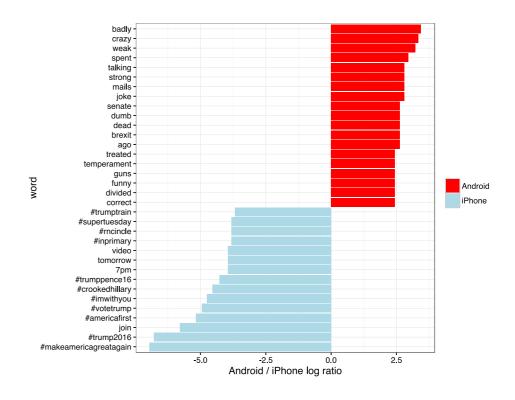
So far, we have not spoken about *randomness* and *stochasticity*. We have, however, spoken about *variation*.

spread in a dataset refers to the fact that in a population of entities there is naturally occuring variation in measurements



### Variation, randomness and stochasticity

Another example: in sets of tweets there is natural variation in the frequency of word usage.



### Variation, randomness and stochasticity

In summary, we can discuss the notion of *variation* without referring to any randomness, stochasticity or noise.

Because, we **do** want to distinguish, when possible:

- natural occuring variation, vs.
- randomness or stochasticity

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- That's difficult to do for all residents.
- Instead we sample (say by randomly sending Twitter surveys), and
   estimate the average and standard deviation of debt in this population
   from the sample.

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Because there is naturally-occuring variation in this population.

So, a simple question to ask is:

How good are our *estimates* of debt mean and standard deviation from sample of 19-30 year old Marylanders?

Another example: suppose we build a predictive model of loan debt for 19-30 year old Marylanders based on other variables (e.g., sex, income, education, wages, etc.) from our sample.

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How good will this model perform when predicting debt in general?

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- Probability captures stochasticity in the sampling process, while
- we *model* naturally occurring variation in measurements in a population of interest.

#### One final word

The term *population* means

the entire collection of entities we want to model

This could include people, but also images, text, chess positions, etc.

#### Random variables

The basic concept in our discussion of probability is the *random variable*.

Task: is a given tweet was generated by a bot?

Action: Sample a tweet **at random** from the set of all tweets ever written and have a human expert decide if it was generated by a bot or not.

Principle: Denote this as a *binary* random variable  $X \in \{0,1\}$ , with value 1 if the tweet is bot-gerneated and 0 otherwise.

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Why is this a random value? Because it depends on the tweet that was randomly sampled.

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We start with a *probability mass function* p:

a.  $p(X=x) \geq 0$  for all values  $x \in \mathcal{D}$ , and

b. 
$$\sum_{x \in \mathcal{D}} p(X = x) = 1$$

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b. the proportion of bot-generated tweets in the set of "all" tweets is p(X=1).

Example The oracle of TWEET

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Suppose we have a magical oracle and know for a *fact* that 70% of "all" tweets are bot-generated.

In that case 
$$p(X = 1) = .7$$
 and  $p(X = 0) = 1 - .7 = .3$ .

cumulative probability distribution P describes the sum of probability up to a given value:

$$P(x) = \sum_{x'\mathcal{D} ext{ s.t. } x' \leq x} p(X=x')$$

#### Expectation

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How many of those do I *expect* to be bot-generated?

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*Expectation* is a formal concept in probability:

$$\mathbb{E}[X] = \sum_{x \in \mathcal{D}} x p(X = x)$$

What is the expectation of X (a single sample) in our tweet example?

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$$\mathbb{E}[X] = 0 \times p(X = 0) + 1 \times p(X = 1) = 0 \times .3 + 1 \times .7 = .7$$

What is the expected number of bot-generated tweets in a sample of n=100 tweets.

Define 
$$Y = X_1 + X_2 + \cdots + X_{100}$$
.

Then we need  $\mathbb{E}[Y]$ 

# (Discrete) Probability distributions

We have  $X_i = \{0,1\}$  for each of the n=100 tweets

Each obtained by uniformly and *independently* sampling from the set of all tweets.

Then, random variable Y is the number of bot-generated tweets in my sample of n=100 tweets.

# (Discrete) Probability distributions

$$egin{aligned} \mathbb{E}[Y] &= \mathbb{E}[X_1 + X_2 + \cdots + X_{100}] \ &= \mathbb{E}[X_1] + \mathbb{E}[X_2] + \cdots + \mathbb{E}[X_{100}] \ &= .7 + .7 + \cdots + .7 \ &= 100 imes .7 \ &= 70 \end{aligned}$$

# (Discrete) Probability distributions

This uses some facts about expectation you can show in general.

(1) For any pair of random variables  $X_1$  and  $X_2$ ,

$$\mathbb{E}[X_1+X_2]=\mathbb{E}[X_1]+\mathbb{E}[X_2].$$

(2) For any random variable X and constant a,  $\mathbb{E}[aX] = a\mathbb{E}[X]$ .

So far we assume we have access to an oracle that told us p(X=1)=.7.

In reality, we *don't*.

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For our tweet analysis task, we need to *estimate* the proportion of "all" tweets that are bot-generated.

This is where our probability model and the expectation we derive from it comes in.

Given data  $x_1, x_2, x_3, \ldots, x_{100}$ ,

With 67 of those tweets labeled as bot-generated (i.e.,  $x_i=1$  for 67 of them)

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We *expect* y=np with p=p(X=1)

Use that observation to *estimate* p!

$$np=67 \Rightarrow \ 100p=67 \Rightarrow \ \hat{p}=rac{67}{100} \Rightarrow \ \hat{p}=.67$$

Our estimate ( $\frac{p}=.67$ ) is wrong, but close.

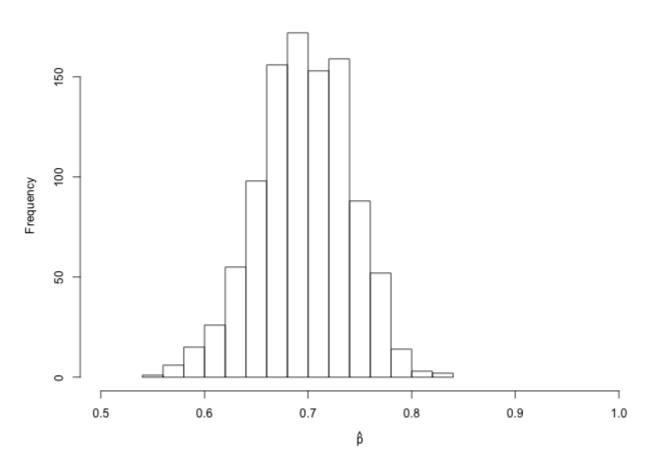
Can we ever get it right?

Can I say how wrong I should expect my estimates to be?

Notice that our estimate of  $\hat{p}$  is the sample *mean* of  $x_1, x_2, \ldots, x_n$ .

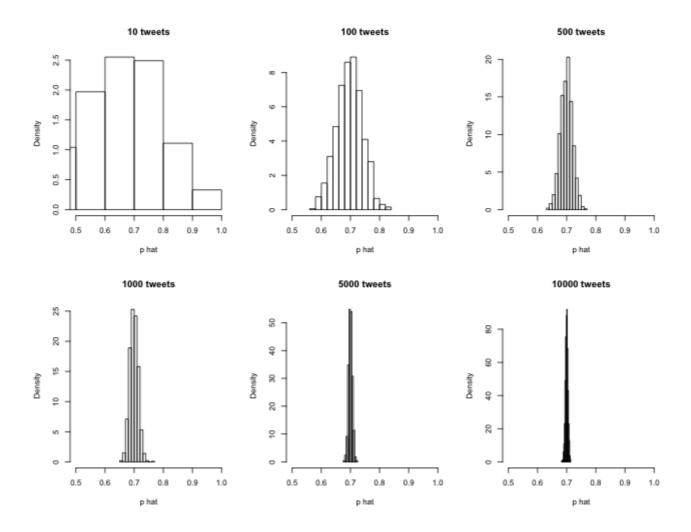
Let's go back to our oracle of tweet to do a thought experiment and replicate how we derived our estimate from 100 tweets a few thousand times.

#### Distribution of p estimates from 100 tweets



What does this say about our estimates of the proportion of botgenerated tweets if we use 100 tweets in our sample?

Now what if instead of sampling n=100 tweets we used other sample sizes?



We can make a couple of observations:

- 1. The distribution of estimate  $\hat{p}$  is  $\emph{centered}$  at p=.7, our unknown  $\emph{population}$  proportion, and
- 2. The *spread* of the distribution **decreases** as the number of samples n increases.

This was a simulation, we faked the data generating procedure.

In reality, we can't.

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In reality, we can't.

What to do we do then?

- (1) Math, or
- (2) Resample

Our simulation is an illustration of two central tenets of statistics:

- (a) The law of large numbers (LLN)
- (b) The central limit theorem (CLT)

Law of large numbers (LLN)

Given *independently* sampled random variables  $X_1, X_2, \cdots, X_n$  with  $\mathbb{E}[X_i] = \mu$  for all i,

$$rac{1}{n}\sum_i X_i o \mu, ext{ as } n o \infty$$

I.E.  $\overline{x}$  tends to the expected value  $\mu$  (under some assumptions beyond the scope of this class) regardless of the distribution  $X_i$ .

Central Limit Theorem (CLT)

The LLN says that estimates built using the sample mean will tend to the correct answer

The CLT describes how these estimates are *spread* around the correct answer.

Here we will use the concept of *variance* which is expected *spread*, measured in squared distance, from the *expected value* of a random variable:

$$\mathrm{var}[\mathrm{X}] = \mathbb{E}[(X - \mathbb{E}[X])^2]$$

$$egin{aligned} ext{var}[ ext{X}] &= \sum_{\mathcal{D}} (x - \mathbb{E}[X])^2 p(X = x) \ &= (0 - p)^2 imes (1 - p) + (1 - p)^2 imes p \ &= p^2 (1 - p) + (1 - p)^2 p \ &= p(1 - p)(p + (1 - p)) \ &= p(1 - p)(p - p + 1) \ &= p(1 - p) \end{aligned}$$

$$P\left(rac{1}{n}\sum_{i=1}X_i
ight)
ightarrow N\left(\mu,rac{\sigma}{n}
ight), ext{ as } n
ightarrow \infty$$

This says, that as sample size n increases, the distribution of sample means is well approximated by a **normal distribution**.

This means we can approximate the *expected error* of our estimates well.

The normal distribution

Random variable  $Y = \sum_{i=1}^{n} X_i$  is *continuous*.

The normal distribution describes the distribution of *continuous* random variables over the range  $(-\infty, \infty)$  using two parameters:

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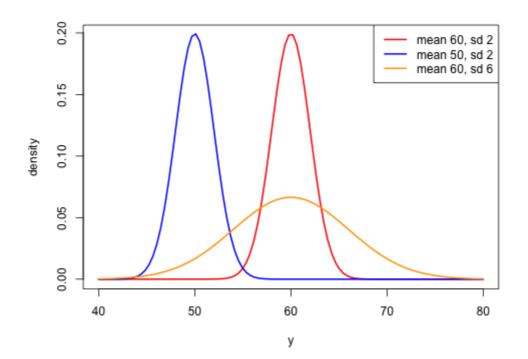
mean  $\mu$  and standard deviation  $\sigma$ .

We write " Y is normally distributed with mean  $\mu$  and standard deviation  $\sigma$ " as  $Y \sim N(\mu, \sigma)$ .

Continuous random variables are described by a *probability density function*. For normally distributed random variables:

$$p(Y=y) = rac{1}{\sqrt{2\pi}\sigma} \exp\left\{-rac{1}{2} \left(rac{y-\mu}{\sigma}
ight)^2
ight\}$$

Three examples of normal probability density functions with mean  $\mu=60,50,60$  and standard deviation  $\sigma=2,2,6$ :



Like the discrete case, probability density functions for continuous random variables need to satisfy certain conditions:

a. 
$$p(Y=y) \geq 0$$
 for all values  $Y \in (-\infty, \infty)$ , and b.  $\int_{-\infty}^{\infty} p(Y=y) dy = 1$ 

One way of interpreting the density function of the normal distribution is that probability decays exponentially with rate  $\sigma$  based on squared distance to the mean  $\mu$ . (Here is squared distance again!)

p(Y=y) \propto \exp \left{ -{\frac{1}{2\sigma^2} (y-\mu)^2} \right }

Also, notice the term inside the square?

$$z=\left(rac{y-\mu}{\sigma}
ight)$$

this is the standardization transformation we saw before.

The name standardization comes from the standard normal distribution N(0,1) (mean 0 and standard deviation 1),

Which is very convenient to work with because it's density function is much simpler:

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In fact, if random variable  $Y \sim N(\mu,\sigma)$  then random variable  $Z = rac{Y-\mu}{\sigma} \sim N(0,1).$ 

#### (Continuous) Random Variables

One more technicality:

The cumulative probability function for continuous random variables is given by

$$P(Y \leq y) = \int_{\mathcal{D}} p(Y=y) dy$$

where  $\mathcal D$  is the range of values random variable Y can take (e.g., for normal distribution  $\mathcal D=(-\infty,\infty)$ )

#### **CLT** continued

We need one last bit of terminology to finish the statement of the CLT.

Consider data  $X_1, X_2, \cdots, X_n$  with  $\mathbb{E}[X_i] = \mu$  for all i, and  $\mathrm{var}(X_i) = \sigma^2$  for all i,

and sample mean  $Y=rac{1}{n}\sum_i X_i$  .

The standard deviation of Y is called the *standard error*.

$$\operatorname{se}(Y) = \frac{\sigma}{\sqrt{n}}$$

Now we can restate the CLT statement precisely:

the distribution of 
$$Y$$
 tends *towards*  $N\left(\mu, \frac{\sigma}{\sqrt{n}}\right)$  as  $n \to \infty$ .

This says, that as sample size increases the distribution of sample means is well approximated by a normal distribution,

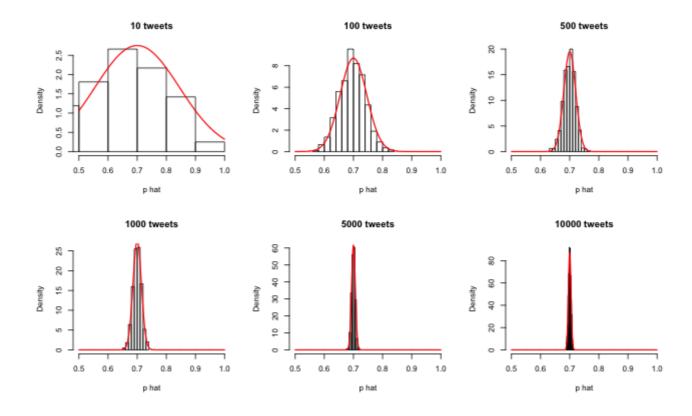
and that the spread of the distribution goes to zero at the rate  $\sqrt{n}$ .

*Disclaimer* There a few mathematical subtleties. Two important ones are that

a.  $X_1, \ldots, X_n$  are iid (independent, identically distributed) random variables, and

b. 
$$\mathrm{var}[X] < \infty$$

Let's redo our simulated replications of our tweet samples to illustrate the CLT at work:



Here we see the three main points of the LLN and CLT:

- (1) the normal density is centered around  $\mu = .7$ ,
- (2) the normal approximation gets better as n increases, and
- (3) the standard error goes to 0 as n increases.

The Bootstrap Procedure

What if the conditions that we used for the CLT don't hold?

For instance, samples  $X_i$  may not be independent. What can we do then, how can we say something about the precision of sample mean estimate Y?

The Bootstrap Procedure

A useful procedure to use in this case is the **bootstrap**.

It is based on using *randomization* to simulate the stochasticity resulting from the population sampling procedure we are trying to capture in our analysis.

The Bootstrap Procedure

The main idea is the following: given observations  $x_1, \ldots, x_n$ 

and the estimate  $y = rac{1}{n} \sum_{i=1}^n x_i$  ,

what can we say about the standard error of y?

The Bootstrap Procedure

There are two challenges here:

- 1) our estimation procedure is deterministic, that is, if I compute the sample mean of a specific dataset, I will always get the same answer; and
- 2) we should retain whatever properties of estimate y result from obtaining it from n samples.

The Bootstrap Procedure

The bootstrap is a randomization procedure that measures the variance of estimate y,

using randomization to address challenge (1),

but doing so with randomized samples of size n, addressing challenge (2).

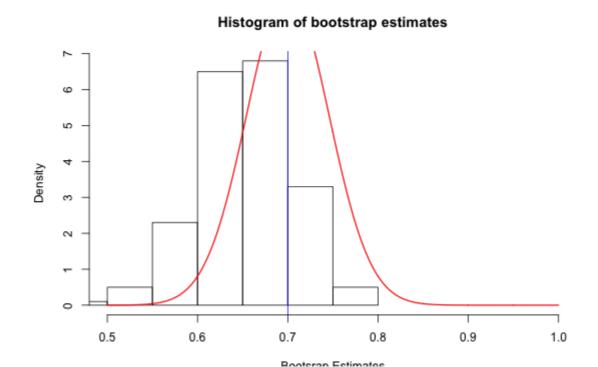
The Bootstrap Procedure

The procedure goes as follows:

- 1. Generate B random datasets by sampling with replacement from dataset  $x_1, \ldots, x_n$ . Denote randomized dataset b as  $x_{1b}, \ldots, x_{nb}$ .
- 2. Construct estimates from *each* dataset,  $y_b = \frac{1}{n} \sum_i x_{ib}$
- 3. Compute center (mean) and spread (variance) of estimates  $y_b$

The Bootstrap Procedure

Let's see how this works on tweet oracle example



The Bootstrap Procedure

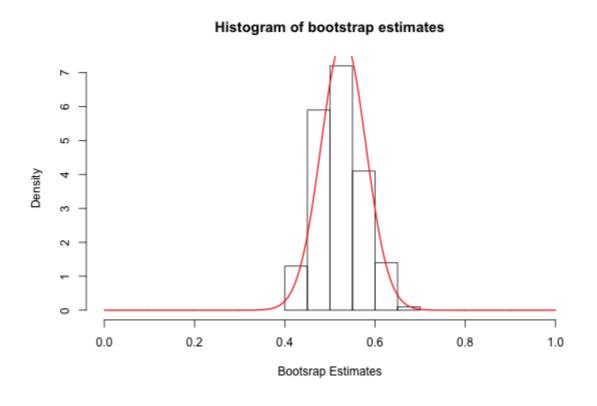
Not great, math works better when conditions are met.

The Bootstrap Procedure

Let's look at a case where we don't expect the normal approximation to not work so well by making samples not identically distributed.

Let's make a new ORACLE of tweet where the probability of a tweet being bot-generated depends on the previous tweet

#### The Bootstrap Procedure



The Bootstrap Procedure

Here, an analysis based on the classical CLT is not appropriate (  $X_i$  s are not independent)

But the bootstrap analysis gives some information about the variability of our estimates.